

Combinatorial Markov chains

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$$\mathbb{F}_n := \{x \in \mathbb{F} : \phi(x) = n\}$$

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- A **combinatorial Markov chain** (CMC) is a Markov chain $X = (X_n)_{n \in \mathbb{N}}$ that is adapted to \mathbb{F} :

$$P(X_n \in \mathbb{F}_n) = 1 \quad \text{for all } n \in \mathbb{N}.$$

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AofA context: Consider an sequential algorithm that transforms a sequence $(\eta_n)_{n \in \mathbb{N}}$ of input values into a sequence $(x_n)_{n \in \mathbb{N}}$ of combinatorial objects such that

$$x_n \in \mathbb{F}_n, \quad x_{n+1} \text{ depends on } x_n \text{ and } \eta_{n+1} \text{ only.}$$

For i.i.d. input this algorithm generates a CMC.

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$$E(\mathbb{F}) = \{(x, y) : x, y \in \mathbb{F}, x \hookrightarrow y\}.$$

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More assumptions:

- **weak connectedness** – there is a path from the root to each node,

$$\forall x \in \mathbb{F} \exists y_i \in \mathbb{F}_i, i = 1, \dots, n := \phi(x) : \\ y_1 = e, y_n = x, (y_{i-1}, y_i) \in E(\mathbb{F}) \text{ for } i = 2, \dots, n$$

- the **transition probabilities** $p(x, y) = P(X_{n+1} = y | X_n = x)$ are adapted to this structure,

$$p(x, y) > 0 \iff x \hookrightarrow y,$$

Example 1: Random walk, urns, and special numbers

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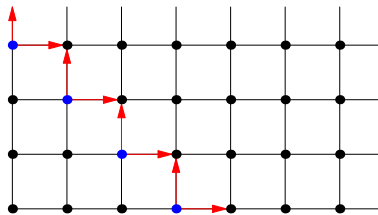
$$\mathbb{F} = \mathbb{N}_0 \times \mathbb{N}_0, \quad \phi(i, j) = i + j + 1,$$

$$(i, j) \hookrightarrow (i + 1, j), \quad (i, j) \hookrightarrow (i, j + 1)$$

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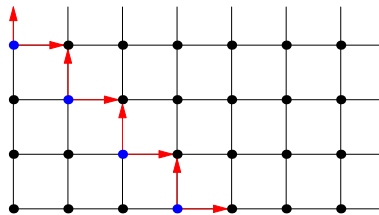
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CMC	$p((i, j), (i + 1, j))$	$\mathcal{L}(X_n)$
north-east random walk	param. $\theta \in (0, 1)$	$\binom{i+j}{i} \theta^i (1 - \theta)^j$
"	$\theta = 1/2$	\sim binomial coefficients
record walk	$1/(i + j + 2)$	\sim Stirling(1) numbers
Friedman urn	$(j + 1)/(i + j + 2)$	\sim Euler numbers
Pólya urn	$(i + 1)/(i + j + 2)$	uniform distribution

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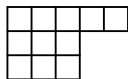
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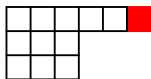
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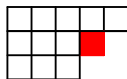
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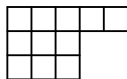
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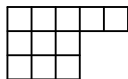
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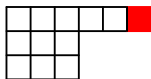
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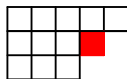
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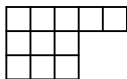
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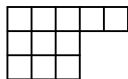
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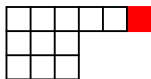
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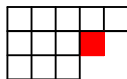
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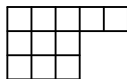
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- **Thoma:** $p(\lambda, \mu) = H(\lambda)/H(\mu)$, hook length formula.
- The **restaurant (or cycle) chain:**
 - If $\mu_j = \lambda_j + 1$, $p(\lambda, \mu) = \lambda_j \cdot \#\{i \geq j : \lambda_i = \lambda_j\} / (n + 1)$.
 - If $\mu = (\lambda_1, \dots, \lambda_k, 1)$: $p(\lambda, \mu) = 1 / (n + 1)$

Example 3: Binary trees

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Let $\mathbb{V} = \{0, 1\}^*$ be the set of nodes, with prefix relation ' \leq '.

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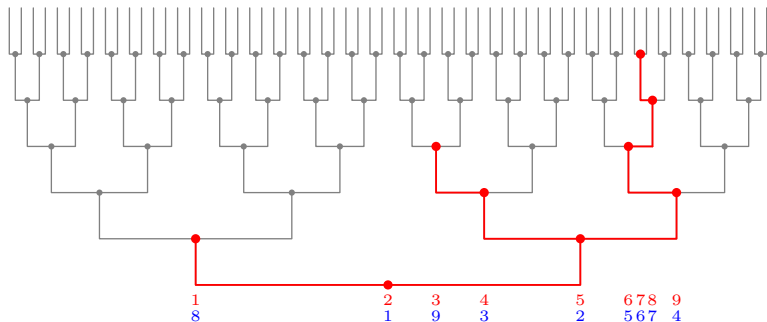
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Connection to representation theory, random walks on groups:

Vershik, Kaimanovich, Kerov (book), Gnedin, Okounkov, Woess ...

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Then $\partial\mathbb{F} := \bar{\mathbb{F}} \setminus \mathbb{F}$ is the **Martin boundary**.

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With $X_1^h = e$ and transitions p^h we obtain another CMC $X^h = (X_n^h)_{n \in \mathbb{N}}$, the **h -transform** of X .

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(reprDM)

The extensions $x \mapsto K(x, \alpha)$, $\alpha \in \partial\mathbb{F}$, are minimal harmonic, and all $h \in \mathbb{H}_+$ with $h(e) = 1$ are mixtures of these:

$$h(x) = \int_{\partial\mathbb{F}} K(x, \alpha) \nu_h(d\alpha)$$

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(struct)

The conditional distribution $\mathcal{L}(X|X_\infty = \alpha)$ is equal to the distribution of the h -transform $\mathcal{L}(X^h)$ with $h = K(\cdot, \alpha)$ (in particular, this is the distribution of a Markov chain).

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Hence, for each $n \in \mathbb{N}$, $(Y_k^n, \mathcal{F}_k^n)_{k=1, \dots, n}$ with

$$Y_k^n := K(x, X_{n+1-k}), \quad \mathcal{F}_k^n = \sigma(X_{n+1-k}, \dots, X_n)$$

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is a **supermartingale**. For the number $U_n(a, b)$ of upcrossings of some interval $[a, b]$ this means

$$EU_n[a, b] \leq \frac{1}{b-a} E(Y_n^n - a)^- = \frac{1}{b-a} (K(x, e) - a)^-.$$

It follows that $(K(x, X_n))_{n \in \mathbb{N}}$ converges almost surely.

Poisson boundary

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A measurable space (E, \mathcal{B}) together with a family ν_x , $x \in \mathbb{F}$, of probability measures on (E, \mathcal{B}) such that $\nu_x \ll \nu_e$ for all $x \in \mathbb{F}$ 'is' the **Poisson boundary** of (X, \mathbb{F}) if

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For each $\phi \in L^\infty$

(reprP)
$$h(x) = \int_E \phi(y) \nu_x(dy), \quad x \in \mathbb{F},$$

is in \mathbb{H}_b , and for each $h \in \mathbb{H}_b$ there is a unique $\phi \in L^\infty$ such that this representation holds.

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For each $h \in \mathbb{H}_b^{\text{st}}$

$$\left(h(n, X_n), \sigma(\{X_m : m \leq n\}) \right)_{n \in \mathbb{N}}$$

is a bounded martingale, hence $h(n, X_n) \rightarrow Y(h)$ a.s., with $Y(h)$ \mathcal{T} -measurable.

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is a bounded martingale, hence $h(n, X_n) \rightarrow Y(h)$ a.s., with $Y(h)$ \mathcal{T} -mesurable. This leads to $L^\infty(\Omega, \mathcal{T}, P \upharpoonright \mathcal{T}) \cong \mathbb{H}_b^{\text{st}}$ via

(reprT)
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- CMCs have the **space-time property**, i.e. $n = \phi(X_n)$, hence tail boundary and Poisson boundary coincide.

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Note: The Poisson boundary for the NE random walk is trivial.

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Boundary theory approach:

- Try to guess Φ ,
- work out Ψ_n ,
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- (5) BTA may lead to interesting functionals (**silhouette** \rightarrow **metric silhouette**).

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Régnier 1989, Rösler 1991, ..., Bindjeme and Fill 2012, Gr. 2012

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Proof uses the contraction method (**Rösler, Rüschemdorf, Neininger**).

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Profiles (from Régnier to Jabbour-Hattab)

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BTA 'explains' the martingale and gives an interpretation of the limit, but does not (easily) lead to a new proof of the known results.

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Also: A representation may lead to convergence in **stronger topologies**.

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Note: The first one is local, the others are global.

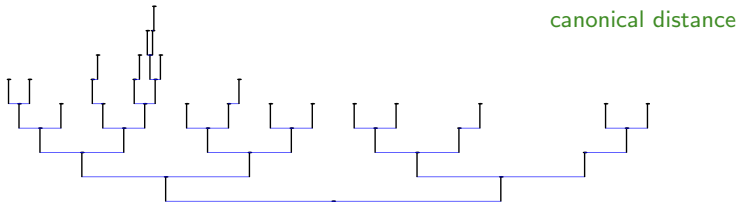
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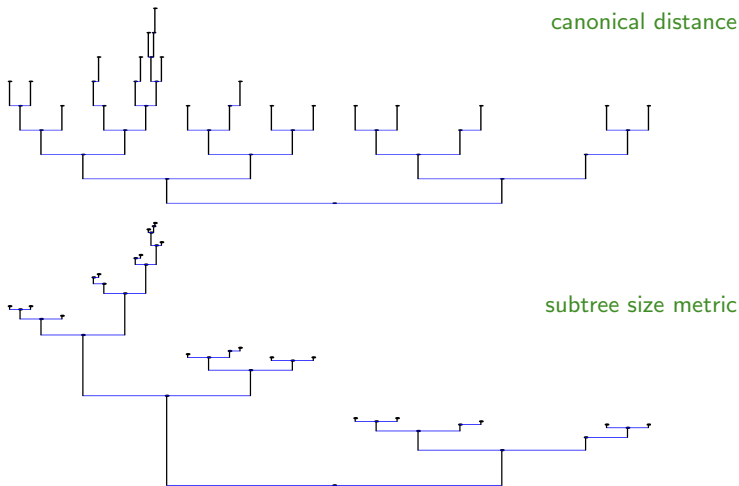
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In view of $X_n \uparrow \mathbb{V}$ we can work with

$$(x_n, d_n) \rightarrow (\mathbb{V}, d_\infty)$$

meaning that

$$\lim_{n \rightarrow \infty} d_n(u, v) = d_\infty(u, v) \quad \text{for all } u, v \in \mathbb{V}.$$

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