

Urn models with multiple drawings

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- 1 Urn models - Introduction
- 2 Urn models with multiple drawings
- 3 Analysis using Analytic Combinatorics

Urn models

- Introduction

Urn models

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Pólya-Eggenberger urns

Urn contains n white and m black balls.

- Every discrete time steps a ball is drawn at random:

$$P_{\text{white}} = \frac{n}{n+m}, P_{\text{black}} = \frac{m}{n+m}.$$

- Color inspection:

White - a white and b black balls are added/removed; Black - c white and d black balls are added/removed; $a, b, c, d \in \mathbb{Z}$.

- 2×2 ball replacement matrix

$$M = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$$

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1 Tenable urns

The process of drawing and adding/removing balls can be continued **ad infinitum**. We start with $W_0 \in \mathbb{N}_0$ white and $B_0 \in \mathbb{N}_0$ black balls: configuration W_n, B_n after n draws?

2 Diminishing urns

The process of drawing and adding/removing balls stops after a **finite** number of steps. We start with $n \in \mathbb{N}_0$ white and $m \in \mathbb{N}_0$ black balls, define so-called absorbing states \mathcal{A} : what is the probability of reaching a state $\alpha \in \mathcal{A}$?

3 Generalizations

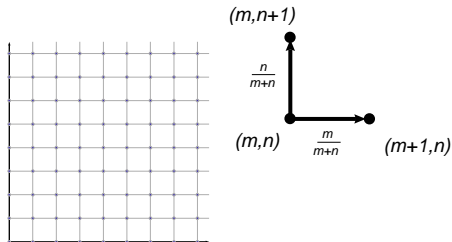
More than two colors.

Pólya-Eggenberger urns

Pólya urn (tenable urns)

$$M = \begin{pmatrix} 1 & 0 \\ 0 & 1 \end{pmatrix}$$

Urn contains n white and m black balls:

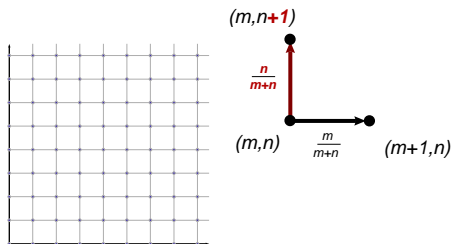


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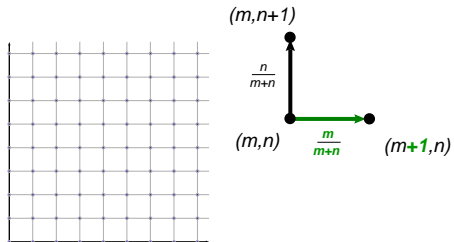


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Pólya-Eggenberger urns

2000-today: several approaches - very interesting developments

Analytic Combinatorics (Symbolic methods, generating functions, etc.)

FLAJOLET, GABARRÓ AND PEKARI; BRENNAN AND PRODINGER; STADJE;
DUMAS, FLAJOLET AND PUYHAUBERT; HWANG, K. AND PANHOLZER; FLAJOLET
AND MORCRETTE; MAHMOUD AND MORCRETTE; MORCRETTE; . . .

Probabilistic methods (stochastic processes, martingales)

KINGMAN²; KINGMAN AND VOLKOV; MAHMOUD^x; PITTEL; JANSON²; CHAUVIN,
POUYANNE ET AL.³; CHEN AND WEI; RENLUND; . . .

Contraction method (for balanced urns)

NEININGER AND KNAPE; CHAUVIN, POUYANNE AND MAILLER; . . .

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An urn model $M = \begin{pmatrix} a & b \\ c & d \end{pmatrix}$ is called **balanced**: $a + b = c + d = \sigma$.

Consequently: $T_n = W_n + B_n = T_0 + n \cdot \sigma$.

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Urn models

- Multiple drawings

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Urn models with multiple drawings

Previously: Urn contains w white and b black balls.

Draw at random a **single** ball: $p_{\text{white}} = \frac{w}{w+b}$, $p_{\text{black}} = \frac{b}{w+b}$.

New model: We draw $m \geq 1$ balls **without replacement**,

$$\mathbb{P}\{k \text{ times white, } (m-k) \text{ times black}\} = \frac{1}{(b+w)^{\underline{m}}} \binom{m}{k} w^k b^{m-k},$$

with $x^{\underline{s}} = x(x-1)\dots(x-s+1)$. Depending on the drawn multiset of white/black balls we add/remove balls.

CHEN AND WEI 2005: Generalized Pólya urn

MAHMOUD 2008: Tenable balanced linear urns ($m = 2$).

RENLUND 2010: Stochastic approximation for tenable urns.

Pólya urn: $M = \begin{pmatrix} c & 0 \\ 0 & c \end{pmatrix}$, $c \in \mathbb{N}$.

Generalization: If we draw $\{W^k S^{m-k}\}$ we add $k \cdot c$ white and $(m - k) \cdot c$ black balls, with $c \in \mathbb{N}$. $(m + 1) \times 2$ -matrix

$$M = \begin{pmatrix} mc & 0 \\ (m-1)c & c \\ \dots & \dots \\ c & (m-1)c \\ 0 & mc \end{pmatrix}$$

Urn is balanced $T_n = W_n + B_n = nmc + T_0$.

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Friedman urn: $M = \begin{pmatrix} 0 & c \\ c & 0 \end{pmatrix}$, $c \in \mathbb{N}$.

Generalization: If we draw $\{W^k S^{m-k}\}$ we add $(m-k) \cdot c$ white and $k \cdot c$ black balls, with $c \in \mathbb{N}$. $(m+1) \times 2$ -matrix

$$M = \begin{pmatrix} 0 & mc \\ c & (m-1)c \\ \vdots & \vdots \\ (m-1)c & c \\ mc & 0 \end{pmatrix}$$

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Urn models

- Results

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Theorem (CHEN AND WEI 2005)

For the generalized Pólya urn the number of white balls W_n after n draws satisfies

$$W_n = \frac{W_n}{T_n} \xrightarrow{\text{a. s.}} W_\infty;$$

W_∞ is absolutely continuous.

Questions (CHEN AND WEI): (1) Is W_∞ beta-distributed? (2)
Explicit results concerning W_n and W_∞ ;

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Urn models with multiple drawings

Theorem (CHEN AND K.)

The expectation and the variance of W_n are given by

$$\mathbb{E}(W_n) = \frac{W_0}{T_0} (nmc + T_0), \text{ and } \mathbb{V}(W_n) = \mathbb{E}(W_n^2) - \mathbb{E}(W_n)^2 \text{ via}$$

$$\mathbb{E}(W_n^2) = \frac{\binom{n-1+\lambda_1}{n} \binom{n-1+\lambda_2}{n}}{\binom{n-1+\frac{T_0}{mc}}{n} \binom{n-1+\frac{T_0-1}{mc}}{n}} \left(W_0^2 + \frac{W_0 c^2 m}{T_0} \sum_{\ell=0}^{n-1} \frac{\ell + \frac{T_0-m}{mc}}{\ell + \frac{T_0-1}{mc}} \frac{\binom{\ell+\frac{T_0}{mc}}{\ell+1} \binom{\ell+\frac{T_0-1}{mc}}{\ell+1}}{\binom{\ell+\lambda_1}{\ell+1} \binom{\ell+\lambda_2}{\ell+1}} \right),$$

with λ_1, λ_2 given by

$$\lambda_{1,2} = \frac{-\frac{1}{2} + mc + T_0 \pm \frac{1}{2} \sqrt{1 + 4mc(1+c)}}{mc}.$$

Urn models with multiple drawings

Theorem (CHEN AND K.)

For $s \geq 1$ the moment $\mathbb{E}(W_n^s)$ is given by

$$\mathbb{E}(W_n^s) = \left(\prod_{j=0}^{n-1} \alpha_{j,s} \right) \cdot \left(W_0^s + \sum_{\ell=0}^{n-1} \frac{\beta_{\ell,s}}{\prod_{j=0}^{\ell} \alpha_{j,s}} \right),$$

with $\alpha_{n,s}, \beta_{n,s}$ determined by

$$\alpha_{n,s} = \sum_{\ell=0}^s c^\ell \frac{\binom{s}{\ell} \binom{m}{\ell}}{\binom{T_n}{\ell}},$$

$$\beta_{n,s} = \sum_{i=2}^s \mathbb{E}(W_n^{s+1-i}) \sum_{\ell=i}^s \binom{s}{\ell} c^\ell \sum_{j=\ell+1-i}^{\ell} (-1)^{j+i-\ell-1} \frac{\left[\begin{smallmatrix} \ell \\ j \end{smallmatrix} \right] \left\{ \begin{smallmatrix} j \\ \ell+1-i \end{smallmatrix} \right\} \binom{m}{j}}{\binom{T_n}{j}}.$$

Here $\left[\begin{smallmatrix} n \\ k \end{smallmatrix} \right]$ and $\left\{ \begin{smallmatrix} n \\ k \end{smallmatrix} \right\}$ denote the Stirling numbers of the first and second kind, respectively.

Proof:

$$W_n \stackrel{(d)}{=} W_{n-1} + c \cdot \xi_n, \quad \xi_n \stackrel{(d)}{=} \text{Hypergeometric}(m, W_{n-1}, B_{n-1});$$

in other words:

$$\mathbb{P}(\xi_n = k \mid \mathcal{F}_{n-1}) = \frac{\binom{W_{n-1}}{k} \binom{T_{n-1} - W_{n-1}}{m-k}}{\binom{T_{n-1}}{m}}.$$

Hence,

$$\mathbb{E}(W_n^s \mid \mathcal{F}_{n-1}) = W_{n-1}^s + \sum_{\ell=1}^s \binom{s}{\ell} W_{n-1}^{s-\ell} c^\ell \sum_{k=1}^m k^\ell \cdot \frac{\binom{W_{n-1}}{k} \binom{T_{n-1} - W_{n-1}}{m-k}}{\binom{T_{n-1}}{m}},$$

By Vandermonde's identity we obtain a recurrence

$$\mathbb{E}(W_n^s) = \alpha_{n-1,s} \cdot \mathbb{E}(W_{n-1}^s) + \beta_{n-1,s}, \quad n, s \geq 1, \quad (1)$$

which leads to the stated result.

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$$\text{expand } k^\ell = \sum_{j=1}^{\ell} \{\ell\}_j k^j.$$

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which leads to the stated result.

Corollary (CHEN AND K.)

The limits $\lim_{n \rightarrow \infty} \mathbb{E}(W_n^s / n^s)$ exist and can be recursively calculated: $\lim_{n \rightarrow \infty} \frac{\mathbb{E}(W_n)}{n} = \frac{W_0 mc}{T_0}$, and

$$\lim_{n \rightarrow \infty} \frac{\mathbb{E}(W_n^2)}{n^2} = \frac{\Gamma(\frac{T_0}{mc})\Gamma(\frac{T_0-1}{mc})}{\Gamma(\lambda_1)\Gamma(\lambda_2)} \times \left(W_0^2 + \frac{W_0 c^2 m}{T_0} \sum_{\ell=0}^{\infty} \frac{\ell + \frac{T_0-m}{mc}}{\ell + \frac{T_0-1}{mc}} \frac{\binom{\ell + \frac{T_0}{mc}}{\ell+1} \binom{\ell + \frac{T_0-1}{mc}}{\ell+1}}{\binom{\ell + \lambda_1}{\ell+1} \binom{\ell + \lambda_2}{\ell+1}} \right).$$

Another model

(CHEN AND K.): draw $m \geq 1$ balls **with replacement**,

$$\mathbb{P}\{k \text{ times white, } (m-k) \text{ times black}\} = \frac{1}{(b+w)^m} \binom{m}{k} w^k b^{m-k}$$

Sampling scheme **influences** the (limiting) distribution.

Theorem

The expected value coincides: $\mathbb{E}(W_n) = \frac{W_0}{T_0}(nmc + T_0)$; but for the second moment

$$\mathbb{E}(W_n^2) = \frac{\binom{n-1+\mu_1}{n} \binom{n-1+\mu_2}{n}}{\left(\binom{n-1+\frac{T_0}{mc}}{n}\right)^2} \left(W_0^2 + \frac{W_0 c^2 m}{T_0} \sum_{\ell=0}^{n-1} \frac{\binom{\ell+\frac{T_0}{mc}}{\ell+1}^2}{\binom{\ell+\mu_1}{\ell+1} \binom{\ell+\mu_2}{\ell+1}} \right)$$

where the values μ_1, μ_2 are given by $\mu_{1,2} = \frac{T_0 + mc \pm c\sqrt{m}}{mc}$.

Urn models with multiple drawings

Friedman urn: $M = \begin{pmatrix} 0 & mc \\ \vdots & \vdots \\ mc & 0 \end{pmatrix}$ We obtain for both

sampling schemes a unified result.

Theorem (MAHMOUD, PANHOLZER AND K.)

The number of white balls W_n after n draws satisfies

$$\frac{W_n}{n} \xrightarrow{\text{(a.s.)}} \frac{cm}{2}, \quad n \rightarrow \infty.$$

Furthermore,

$$\frac{W_n - \frac{1}{2}cmn - \frac{1}{2}T_0}{\sqrt{n}} \xrightarrow{\text{(d)}} \mathcal{N}\left(0, \frac{1}{12}c^2m\right),$$

with convergence of all moments.

Urn models

- Multiple drawings

Analytic combinatorics

Urn models

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Analytic combinatorics

General model:

$$M = \begin{pmatrix} a_0 & b_0 \\ a_1 & b_1 \\ \dots & \dots \\ a_{m-1} & b_{m-1} \\ a_m & b_m \end{pmatrix}.$$

Draw a multiset $\{W^k S^{m-k}\}$: we add a_k white and b_k black balls.

Case $m = 1$: draw a **single** ball

- 2005 FLAJOLET, GABARRÓ AND PEKARI: **first order PDE** for balanced urns
- 2006 DUMAS, FLAJOLET AND PUYHAUBERT: **differential systems** for balanced urns
- 2010-2011 FLAJOLET AND MORCRETTE: **AC** for **unbalanced** urns
- 2013 MORCRETTE: first order PDE for **unbalanced** urns!

Case $m = 2$:

- July 2010 FLAJOLET Second order PDE for Bernoulli-Laplace urn

Analysis using Analytic combinatorics

Balanced urns: $M = \begin{pmatrix} a_0 & b_0 \\ a_1 & b_1 \\ \dots & \dots \\ a_{m-1} & b_{m-1} \\ a_m & b_m \end{pmatrix}$, $\sigma = a_i + b_i$. Approach

of DUMAS, FLAJOLET AND PUYHAUBERT: Study **urn histories** using differential operators:

∂_z : differential operator with respect to z ; $\Theta_z = z \cdot \partial_z$.

Assume we have w white and b black balls $\implies x^w y^b$:

$$y^{m-k} \partial_y^{m-k} x^k \partial_x^k (x^w y^b) = \frac{x^w y^b}{(b+w)^m} \binom{m}{k} w^k b^{m-k},$$

$$\Theta_x^k \Theta_y^{m-k} (x^w y^b) = \frac{x^w y^b}{(b+w)^m} \binom{m}{k} w^k b^{m-k}.$$

Analysis using Analytic combinatorics

We introduce the differential operators

$$\mathcal{D}_M = \sum_{k=0}^m \binom{m}{k} x^{a_k+k} y^{b_k+m-k} x^k y^{m-k} \partial_x^k \partial_y^{m-k},$$

and

$$\mathcal{D}_R = \sum_{k=0}^m \binom{m}{k} x^{a_k} y^{b_k} \Theta_x^k \Theta_y^{m-k}.$$

Proposition

Starting with W_0 white and B_0 black balls the generating function of all urn histories $H(x, y; z) = \sum_{n \geq 0} \mathcal{D}^n x^{W_0} y^{B_0} \frac{z^n}{(n!)^m}$ satisfies

$$\mathcal{D} * H(x, y, z) = \frac{1}{z} \Theta_z^m * H(x, y, z),$$

with $\mathcal{D} = \mathcal{D}_M$ (without replacement) or $\mathcal{D} = \mathcal{D}_R$ (with replacement).

Further simplifications using $\Theta_x + \Theta_y = W_0 + B_0 + \sigma \Theta_z$.

Outlook

- Similar results for diminishing urn models; some second order PDEs are explicitly solvable (reduction to first order):

$$M = \begin{pmatrix} -1 & 0 \\ -1 & 0 \\ 0 & -1 \end{pmatrix}, M = \begin{pmatrix} -1 & 0 \\ -1 & -1 \\ 0 & -1 \end{pmatrix}, M = \begin{pmatrix} -1 & 0 \\ -1 & 0 \\ 1 & -2 \end{pmatrix}, M = \begin{pmatrix} -1 & 0 \\ 0 & -1 \\ 1 & -2 \end{pmatrix},$$
$$M = \begin{pmatrix} -2 & 0 \\ -1 & 0 \\ 0 & -2 \end{pmatrix}, M = \begin{pmatrix} 1 & 0 \\ 1 & 0 \\ 0 & 1 \end{pmatrix} \dots$$

explicit results, moments, limit laws, ...

- Approach of **Morcrette** \Rightarrow higher order PDE for **unbalanced** urns with **multiple drawings**...

Solvable higher order PDEs stemming from urn models

Thanks for your **attention!**

Thanks for your **attention!**